

## Augmenting our AR(4) Model of Inflation

- Adding lagged unemployment to our model of inflationary change, we get:

$$\widehat{\Delta Inf}_t = 1.28 \quad -(0.31)\Delta Inf_{t-1} \quad -(0.39)\Delta Inf_{t-2} \quad +(0.09)\Delta Inf_{t-3} \\ (0.53) \quad (0.09) \quad (0.09) \quad (0.08) \\ -(0.08)\Delta Inf_{t-4} \quad -(0.21)Unemp_{t-1} \\ (0.09) \quad (0.09)$$

- The  $\bar{R}^2 = 0.21$ .
- The corresponding forecast for 2005:I is 3.9%, with a forecast error of -1.5%
- Stock and Watson consider including 2nd through 4th lags for unemployment as well and find the additional regressors are jointly significant at a 1% level.

## The Autoregressive Distributed Lag (ADL) Model

- The models from the previous slide are known as **Autoregressive Distributed Lag (ADL) Model**.
- Specifically, with
  - $p$  lags of the dependent variable and
  - $q$  lags of the additional explanatory variable
 the model is called an **ADP(p,q)** model.
- The unemployment model on the previous slide becomes an ADL(4,1).
- Formally, the model becomes:

$$Y_t = \beta_0 + \beta_1 Y_{t-1} + \cdots + \beta_p Y_{t-p} \\ + \delta_1 X_{t-1} + \cdots + \delta_q X_{t-q} + u_t$$

- An underlying assumption justifying OLS in this case is that:

$$E(u_t | Y_{t-1}, Y_{t-2}, \cdots, X_{t-1}, X_{t-2}, \cdots) = 0 \quad (13)$$

- One can obviously add multiple regressors with differing lags.

## Stationarity

- A key assumption in most time series models is that of **stationarity**:

A time series  $Y_t$  is **stationary** if its probability distribution does not change over time; i.e., if the joint distribution of  $(Y_{s+1}, Y_{s+2}, \dots, Y_{s+T})$  does not depend upon  $s$ .

- Intuitively, since we are wanting to use past values of  $Y_t$  to predict future values of  $Y_t$ , it helps if the historical distribution for  $Y_t$  carries into the future.

## The Key Assumptions Underlying OLS in the AR(p,q) model

Our more general model becomes:

$$\begin{aligned} Y_t &= \beta_0 + \beta_1 Y_{t-1} + \dots + \beta_p Y_{t-p} \\ &+ \delta_{11} X_{1,t-1} + \dots + \delta_{1q_1} X_{1,t-q} \\ &+ \delta_{k1} X_{k,t-1} + \dots + \delta_{kq_k} X_{k,t-q} + u_t \end{aligned}$$

The assumptions underlying OLS are:

- 1  $E(u_t | Y_{t-1}, Y_{t-2}, \dots, X_{1,t-1}, X_{1,t-2}, \dots, X_{k,t-1}, X_{k,t-2}, \dots) = 0$ .
- 2 Has two parts:
  - $(Y_t, X_{1,t}, \dots, X_{k,t})$  has a stationary distribution.
  - $(Y_t, X_{1,t}, \dots, X_{k,t})$  and  $(Y_{t-j}, X_{1,t-j}, \dots, X_{k,t-j})$  become independent at  $j$  gets large (known as **weak dependence**).
- 3  $(Y_t, X_{1,t}, \dots, X_{k,t})$  have nonzero and finite fourth moments.
- 4 There is no perfect multicollinearity.

## Granger Causality

- With the  $ADL(p, q)$  model, we are using lagged values of one variable (i.e., the  $X_t$ 's) to aid in predicting another (i.e., the  $Y_t$ 's).
- The **Granger causality test** essentially is testing the validity of this assumption using an F-test, with the null hypothesis being:  
 $H_0 : \delta_{k1} = \dots = \delta_{kq_k} = 0.$
- As the book notes, the terminology here is different from that used in the previous chapter. We are not saying that lagged values of the  $X_t$ 's cause changes in the  $Y_t$ 's.
- Instead, we are saying that "... past values of the  $X_t$ 's are useful in predicting  $Y_t$ , beyond the information contained in the lagged values of  $Y_t$ .

## Forecast Uncertainty and Forecast Intervals

- Consider the simple  $ADL(1, 1)$  model with a forecasted value for  $Y_{T+1}$  given by

$$\hat{Y}_{T+1|T} = \hat{\beta}_0 + \hat{\beta}_1 Y_T + \hat{\delta}_1 X_T. \quad (14)$$

- The corresponding forecast error is given by

$$\begin{aligned} \text{Forecast error} &= Y_{T+1} - \hat{Y}_{T+1|T} \\ &= [\beta_0 + \beta_1 Y_T + \delta_1 X_T + u_{T+1}] \\ &\quad - [\hat{\beta}_0 + \hat{\beta}_1 Y_T + \hat{\delta}_1 X_T] \\ &= u_{T+1} + [(\beta_0 - \hat{\beta}_0) \\ &\quad + (\beta_1 - \hat{\beta}_1) Y_T + (\delta_1 - \hat{\delta}_1) X_T] \end{aligned}$$

## The Mean Squared Forecast Error (MSFE)

- The error term  $u_{T+1}$ 
  - has a mean of zero
  - is homoskedastic with variance  $\sigma_u^2$  due to the stationarity assumption
  - and is uncorrelated with the OLS estimator.
- As a result:

$$\begin{aligned} MSFE &= E \left[ \left( Y_{T+1} - \hat{Y}_{T+1|T} \right)^2 \right] \\ &= \sigma_u^2 + \text{var} \left[ \left( \beta_0 - \hat{\beta}_0 \right) \right. \\ &\quad \left. + \left( \beta_1 - \hat{\beta}_1 \right) Y_T + \left( \delta_1 - \hat{\delta}_1 \right) X_T \right] \end{aligned}$$

- The former can be estimated using the square of the SER, while the latter requires computing the variance of a weighted average of the parameter estimates (using “test” in STATA).

## Forecast Intervals

- One period ahead forecast intervals are formed in the usual way,

$$\hat{Y}_{T+1|T} \pm z_\alpha \times \sqrt{MSFE}. \quad (15)$$

- If we are forecasting more than one period ahead, the uncertainty becomes larger.
- Consider forecasting two periods ahead in an AR(1) Model. We know that

$$\begin{aligned} Y_{T+2} &= \beta_0 + \beta_1 Y_{T+1} + u_{T+2} \\ &= \beta_0 + \beta_1 [\beta_0 + \beta_1 Y_T + u_{T+1}] + u_{T+2} \\ &= \beta_0 + \beta_1 \beta_0 + \beta_1^2 Y_T + \beta_1 u_{T+1} + u_{T+2} \end{aligned}$$

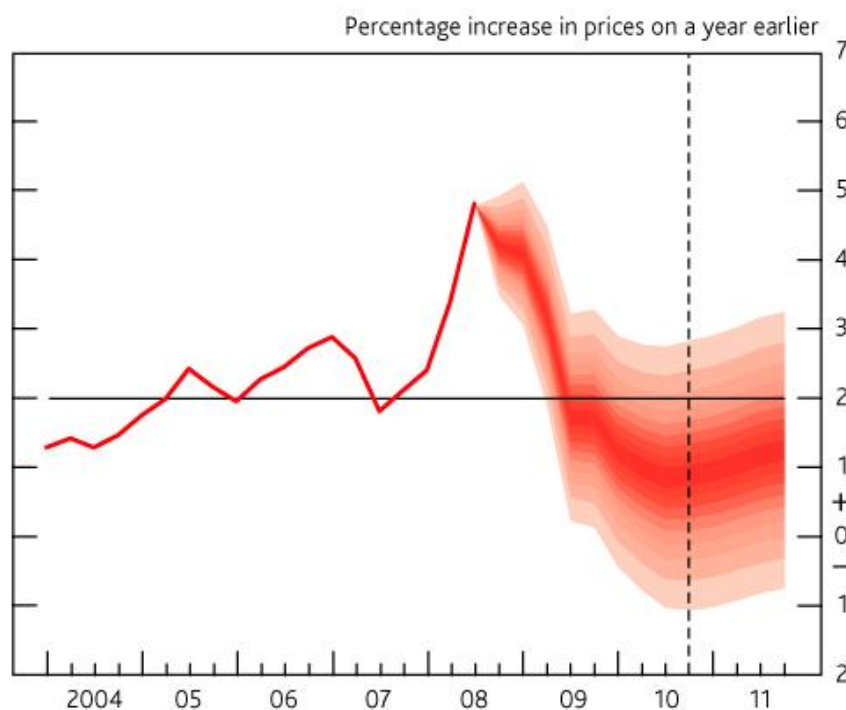
- For a large sample

$$\begin{aligned} RMSFE_2 &= \sqrt{E[(Y_{T+2} - \hat{Y}_{T+2|T})^2]} \\ &\approx \sqrt{E[\beta_1^2 u_{T+1}^2 + u_{T+2}^2]} = \sqrt{(1 + \beta_1^2) \text{var}(u_t)} \approx (1 + \beta_1^2) SER \end{aligned}$$

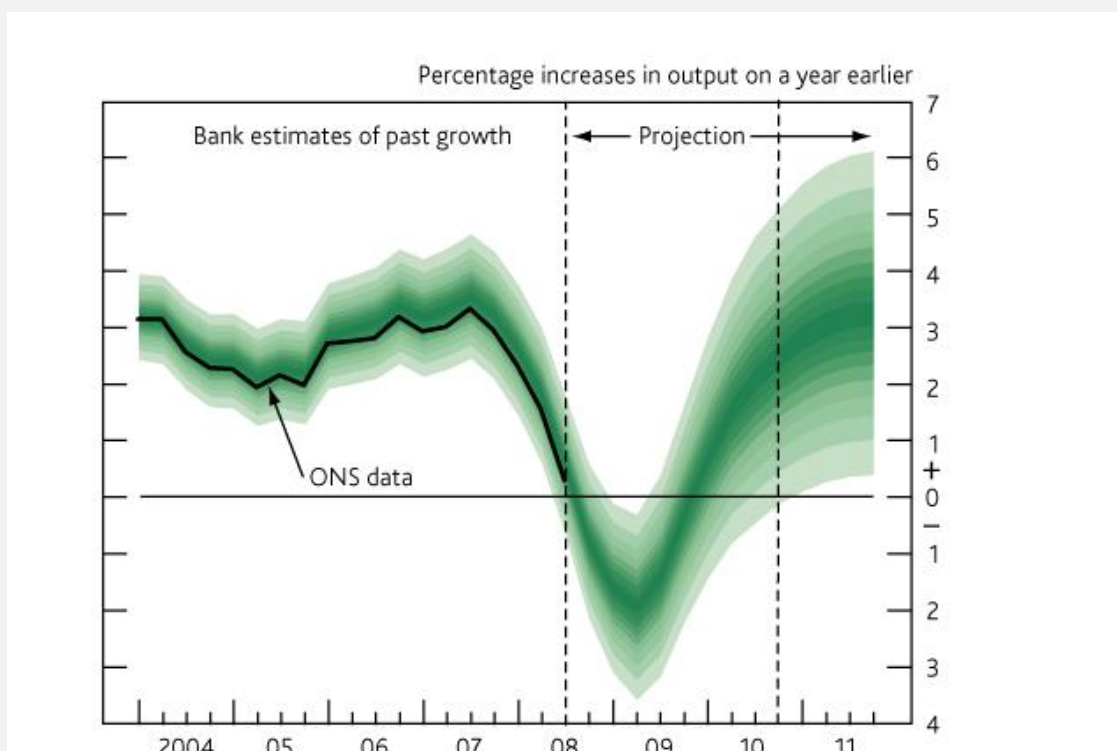
## The River of Blood Graphs

- Stock and Watson note that professional forecasters "...often report confidence intervals that are tighter than 95%..." because otherwise the intervals are of "...limited use in decision making."
- This, of course, is nonsense, since tighter looking 68% intervals (one standard errors) are no more informative, only appearing to be more precise.
- The Bank of England uses a gradated chart, indicating reduced certainty as you consider light shaded regions of the projections (similar to hurricane landing forecasts).

## Example #1: River of Blood Chart for Inflation



## Example #2: River of Blood (Actually Green) Chart for GDP Growth



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### Lag Length Selection

## Selecting Lag Length in an AR(p) Model

- Choosing the order  $p$  of an AR model requires balancing the information lost from omitting useful past lags against estimation error.
- There are a variety of approaches
  - 1 **F-statistic**: Start with a high value of  $p$ , dropping last lag if it is statistically insignificant at a given level.  $p$  - value will be incorrect.
  - 2 **Bayes information criterion (BIC)**:

$$BIC(p) = \ln \left( \frac{SSR(p)}{T} \right) + (p + 1) \frac{\ln(T)}{T} \quad (16)$$

This tradesoff the reduction in SSR versus number of lags.

- 3 **Akaike Information criterion (AIC)**:

$$AIC(p) = \ln \left( \frac{SSR(p)}{T} \right) + (p + 1) \frac{2}{T} \quad (17)$$

AIC tends to overestimate the number of lags.

## BIC for US Inflation Case

**TABLE 14.4** The Bayes Information Criterion (BIC) and the  $R^2$  for Autoregressive Models of U.S. Inflation, 1962–2004

$p$	$SSR(p)/T$	$\ln(SSR(p)/T)$	$(p+1)\ln(T)/T$	$BIC(p)$	$R^2$
0	2.900	1.065	0.030	1.095	0.000
1	2.737	1.007	0.060	1.067	0.056
2	2.375	0.865	0.090	0.955	0.181
3	2.311	0.838	0.120	0.957	0.203
4	2.309	0.837	0.150	0.986	0.204
5	2.308	0.836	0.180	1.016	0.204
6	2.308	0.836	0.209	1.046	0.204

## Lag Length with Multiple Regressors

- A similar set of issues arise when there are multiple regressors, as in a  $ADL(p,q)$  setting.
- Again, an F-statistic approach can be used, but will tend to overstate the various lag lengths.
- There are counterpart expressions for the BIC and AIC. For example, for the BIC, we have

$$BIC(K) = \ln\left(\frac{SSR(p)}{T}\right) + K \frac{\ln(T)}{T} \quad (18)$$

where  $K$  denotes the number of coefficients including the intercept.

## Trends

- A **trend** is a persistent, long-term movement of a variable over time, with the series fluctuating around this trend.
- In our earlier figures, we saw
  - Figure 14.1a: An upward trend in inflation prior to 1982 and a downward trend thereafter.
  - Figure 14.1c: An upward trend in Japanese GDP, or alternatively a downward trend in the growth rate of its GDP.
- Trends can be:
  - **deterministic**: a nonrandom function of time, such as a linear trend equal to  $a \times t$  where  $a$  is a constant.
  - **stochastic**: a random function varying over time. These are typically viewed as more realistic, allowing for fluctuations or “surprises” over time.

## The Random Walk

- The **random walk** is a particular stochastic trend, with

$$Y_t = Y_{t-1} + u_t \quad (19)$$

where the  $u_t$ 's (i.e., the “steps”) are *i.i.d.*.

- In the case of a random walk, the best forecast of tomorrow's value for the variable is today's value.
- A slightly more general version of the random walk is the random walk **with drift**; i.e.,

$$Y_t = \beta_0 + Y_{t-1} + u_t \quad (20)$$

where  $\beta_0$  is the drift term.