

**Econ 573**

Helle Bunzel

**ASYMPTOTIC THEORY – Handout #1****Summary****Definition: Converge almost surely**

A sequence of random variables  $\{X_n\}$  is said to converge almost surely (a.s.) to a random variable  $X$ , denoted by  $X_n \xrightarrow{a.s.} X$ , if

$$\Pr(\lim_{n \rightarrow \infty} X_n = X) = 1$$

An equivalent way of defining almost sure convergence is by

$$\lim_{n \rightarrow \infty} \Pr(|X_m - X| < \epsilon, \text{ all } m \geq n) = 1$$

The almost sure convergence is the mode of convergence associated with the strong law of large numbers (SLLN).

**Definition: Converge in r'th mean**

Let  $\{X_n\}$  be a sequence of random variables such that  $E(|X_n|^r) < \infty$  for all  $n \in N$  and  $E(|X|^r) < \infty$  for  $r > 0$ , then the sequence converges to  $X$  in r'th mean, denoted by  $X_n \rightarrow X$ , if

$$\lim_{n \rightarrow \infty} E(|X_n - X|^r) = 0$$

Of particular interest in what follows is the convergence in mean ( $r = 1$ ) and mean square ( $r = 2$ ).

**Definition: Converge in probability**

A sequence of random variables  $\{X_n\}$  is said to converge in probability to a random variable  $X$ , denoted by  $X_n \xrightarrow{P} X$ , if

$$\lim_{n \rightarrow \infty} \Pr(|X_n - X| < \epsilon) = 1$$

**Definition: Converge in distribution**

A sequence of random variables  $\{X_n\}$  with distribution functions  $\{F_n(x)\}$  is said to converge in distribution to  $X$ , denoted by  $X_n \xrightarrow{D} X$  if

$$\lim_{n \rightarrow \infty} F_n(x) = F(x)$$

at every continuity point  $x$  of  $F(x)$ .

The mode of convergence related to the central limit theorem (CLT) is that of convergence in distribution.

Another way of looking at the convergence in probability is:

Consider the sequence of stochastic variables

$$X_1, X_2, \dots, X_n, \dots$$

and the sequence of their distribution functions

$$F_1, F_2, \dots, F_n, \dots$$

$X_n$  converges to  $X$  in probability if

$$\lim_{n \rightarrow \infty} Pr ( |X_n - X| > \epsilon ) = 0 \text{ for any } \epsilon > 0$$

For a given  $\epsilon$  and the distribution function  $F_n(\cdot)$  we can find the probability  $P(X_n < X - \epsilon) + P(X_n > X + \epsilon)$ .

The sequence of these probabilities is

$$a_1, a_2, \dots, a_n, \dots$$

Convergence in probability requires that this sequence has the limit 0.

Convergence in probability is stated as

$$X_n \xrightarrow{P} X \text{ or } \text{plim } X_n = X$$

**LEMMA**

Let  $\{X_n\}$  be a random vector sequence and  $g(\cdot): \mathbb{R}^K \rightarrow \mathbb{R}$  a continuous function at  $X$ , then

$$1) \quad X_n \xrightarrow{a.s.} X \Rightarrow g(X_n) \xrightarrow{a.s.} g(X)$$

$$2) \quad X_n \xrightarrow{P} X \Rightarrow g(X_n) \xrightarrow{P} g(X)$$

$$3) \quad X_n \xrightarrow{D} X \Rightarrow g(X_n) \xrightarrow{D} g(X)$$

This lemma can be used to prove the following results:

$$\text{plim}(X_n + Y_n) = \text{plim } X_n + \text{plim } Y_n = \mu + ?$$

$$\text{plim}(X_n \cdot Y_n) = \text{plim } X_n \cdot \text{plim } Y_n = \mu \cdot ?$$

$$\text{plim} \left( \frac{X_n}{Y_n} \right) = \frac{\text{plim } X_n}{\text{plim } Y_n} = \frac{\mathbf{m}}{\mathbf{x}}$$

$$\text{plim } g(X_n) = g(\text{plim } X_n) = g(\mu)$$

$$\text{plim } X_n^p = (\text{plim } X_n)^p = \mathbf{m}^p$$

$$\text{plim } A_T^{-1} = (\text{plim } A_T)^{-1}$$

$$\text{plim } A_T \cdot B_T = \text{plim } A_T \cdot \text{plim } B_T$$

**Definition: Consistent estimator**

$\hat{q}_n$  is a consistent estimator of  $q$  if

$$\text{plim } \hat{q}_n = q$$

**Theorem. Sufficient conditions for convergence in probability**

$$\text{If } \lim_{n \rightarrow \infty} E(\hat{q}_n) = q$$

$$\text{and } \lim_{n \rightarrow \infty} V(\hat{q}_n) = 0$$

$$\text{then } \text{plim } \hat{q}_n = q.$$

Otherwise convergence in probability is often proved using Chebyshev's inequality

$$\Pr (|X - \mathbf{m}| \geq k \cdot \mathbf{s}) \leq \frac{1}{k^2}$$

Convergence in distribution can be proved by finding the limit of the sequence of characteristic functions

$$f_n = E(\exp(i \cdot n \cdot X_n))$$

$$f_n \rightarrow f \text{ for } n \rightarrow \infty \Rightarrow F_n \xrightarrow{D} F$$

For simple cases this is not necessary.

**Example**

Let  $X_i \sim \text{nid}(\mu, \sigma^2)$

$$\text{Let } \bar{X}_n = \frac{\sum_{i=1}^n X_i}{n}$$

Then the sequence  $\{X_n\}$

$$\bar{X}_1, \bar{X}_2, \dots, \bar{X}_n, \dots$$

has the following sequence of distribution functions

$$N(\mu, \frac{\sigma}{\sqrt{1}}), N(\mu, \frac{\sigma}{\sqrt{2}}), \dots, N(\mu, \frac{\sigma}{\sqrt{n}})$$

This sequence of distribution functions degenerates to the non-stochastic constant  $\mu$  because the variance goes to zero.

Instead we can look at  $\sqrt{n}\bar{X}$ , or usually we look at

$$\sqrt{n} Z_n = \sqrt{n} \frac{\bar{X} - \mu}{\sigma}$$

$\{\sqrt{n}Z_n\}$  has the following sequence of distribution functions

$$N(0,1), N(0,1), \dots, N(0,1), \dots$$

which obviously has the limit  $N(0,1)$ .

Otherwise one of the following Lemmas, the Cramer's Theorem, or the Mann-Wald Theorem, can be used to find the limiting distribution function.

**Lemma**

Let  $\{X_n, Y_n\}$  be a sequence of pair of random  $k \times 1$  vectors.

Then:

$$(1) \quad \text{If } (X_n - Y_n) \xrightarrow{P} 0 \text{ and } X_n \xrightarrow{D} X \Rightarrow Y_n \xrightarrow{D} X :$$

$$(2) \quad \text{If } X_n \xrightarrow{D} X \text{ and } Y_n \xrightarrow{P} 0 \Rightarrow X_n \cdot Y_n \xrightarrow{P} 0$$

$$(3) \quad \begin{aligned} &\text{If } X_n \xrightarrow{D} X \text{ and } Y_n \xrightarrow{P} C \text{ (constant)} \Rightarrow (X_n + Y_n) \xrightarrow{D} X + C \\ &\text{If } X_n \xrightarrow{D} X \text{ and } Y_n \xrightarrow{P} C \text{ (constant)} \Rightarrow Y_n X_n \rightarrow CX \end{aligned}$$

$$(4) \quad (\text{for } Y_n \text{ and } C \text{ kxk non-singular}) \quad Y_n^{-1} X_n \xrightarrow{D} C^{-1} X$$

### CRAMÉRS THEOREM

Let  $H_n$  be  $k \times m$  matrix, such that

$$(i) \quad \underset{n \rightarrow \infty}{plim} H_n = H$$

where  $H$  is a  $k \times m$  matrix with rank  $k \times m$ .

Let  $\eta_n$  be an  $m$  element stochastic vector such that

$$(ii) \quad \sqrt{n} (\mathbf{x}_n - \mathbf{x}) \xrightarrow{D} N(0, \Sigma)$$

where  $\eta$  is an  $m$  element non-stochastic vector, and  $S$  is an  $m \times m$  positive semidefinite matrix.

Define the  $k$  element stochastic vector  $\eta_n = H_n \eta_n$ . Then

$$(1) \quad \sqrt{n} (\mathbf{h}_n - H \mathbf{x}) \xrightarrow{D} N(0, H \Sigma H')$$

### MANN-WALD THEOREM

Let  $z_n$  be a  $k$  element stochastic vector such that

$$(i) \quad \underset{n}{plim} \frac{1}{n} \sum_{t=1}^n Z_t Z_t' = Q \text{ or } \underset{n}{plim} \frac{Z'Z}{n} = Q$$

where  $Q$  is a  $k \times k$  positive definite matrix.

Assume

$$(ii) \quad u_t \sim \text{iid}(0, s^2)$$

$$(iii) \quad E(Z_t u_t) = 0$$

Then

$$(1) \quad \underset{n}{plim} \frac{1}{n} \sum_{t=1}^n Z_t u_t = 0 \text{ or } \underset{n}{plim} \frac{Z'u}{n} = 0$$

and

$$(2) \quad \sqrt{n} \left( \frac{1}{n} \sum_{t=1}^n Z_t u_t \right) = \frac{1}{\sqrt{n}} \left( \sum_{t=1}^n Z_t u_t \right) \xrightarrow{D} N(0, \mathbf{s}^2 Q)$$

Using a Taylor series approximation and Cramer's Theorem, the following theorem can be proved.

**Theorem. Asymptotic Distribution of a Nonlinear Function**

$$\text{If } \sqrt{n} (\hat{\mathbf{q}} - \mathbf{q}) \xrightarrow{D} N(0, \Sigma)$$

and if  $g(\cdot)$  is a continuous function and its first and second order derivatives exist, then

$$\sqrt{n} (g(\hat{\mathbf{q}}) - g(\mathbf{q})) \xrightarrow{D} N(0, G(\mathbf{q})\Sigma G(\mathbf{q})')$$

$$\text{where } G(\theta) = \left\{ \frac{\partial g(\theta)}{\partial \theta} \right\}$$

**Lemma. plim of sample moment**

Let  $m^k$  be a sample moment of  $\mu_k = E(x^k)$ .  
Assume that  $\mu_{2k}$  exists then  $\text{plim } m_k = \mu_k$ .

**Lindberg-Feller CLT (Greene Thm 4.14)**

Suppose that  $x_1, x_2, x_3, \dots, x_n$  are a sample of random vectors such that  $E[x_i] = \mathbf{m}_i$ ,  $\text{Var}[x_i] = Q_i$ ,

and all mixed third moments of the multivariate distribution are finite. Let

$$\bar{\mathbf{m}}_n = \frac{1}{n} \sum_{i=1}^n \mathbf{m}_i,$$

$$\bar{Q}_n = \frac{1}{n} \sum_{i=1}^n Q_i.$$

We assume that

$$\lim_{n \rightarrow \infty} \bar{Q}_n = Q,$$

where  $Q$  is a finite, positive definite matrix, and that for every  $i$ ,

$$\lim_{n \rightarrow \infty} (\overline{nQ_n})^{-1} Q_i = 0.$$

Then

$$\sqrt{n}(\bar{x}_n - \bar{\mathbf{m}}_n) \xrightarrow{d} N(0, Q)$$