



Causal Effects of Mental Health Conditions on Food Insecurity and the Mitigating Role of SNAP

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Motivation

Mental health conditions affect many individuals:

- 18% of U.S. adults have suffered from mental illness

Studies find association between mental health and food security:

- Depressive symptoms and food insecurity
- Mother's mental health problems and food insecurity

No known research on **causal** impact of mental health on food security that accounts for:

- True mental health status as potentially misreported
- Mitigating role of SNAP in effects of mental illness

Research Objectives

We study **causal effects** of adult mental health conditions on food security and mitigating role of SNAP:

- To what extent does mental health of low-income adults causally affect food security of their families? Do these effects vary across socioeconomic and demographic characteristics?
- Do estimated causal effects differ by whether family participates in SNAP? Does SNAP play a meaningful, mitigating role in these relationships?

Methodological Challenge

Identifying **causal** effect of mental health is difficult:

- **Endogeneity**: same unobservables simultaneously affect food security and mental health
 - Simple regression methods produce **inconsistent** estimates of causal effects
- **Measurement error**: mental disorders are often misdiagnosed, survey instruments have flaws, stigma leads to underreporting
 - Treatment variables are binary → error is **nonclassical**
 - Standard IV methods produce **inconsistent** estimates

Assessing whether SNAP mitigates effect of mental illness on food security is challenging because SNAP participation is **endogenous** and **misreported**

We develop partial identification methodology to quantify **joint effect** on food security of **two** potentially mismeasured, endogenous treatments: mental illness and SNAP participation

Main Data Source

National Health Interview Survey (**NHIS**):

- Principal source of info on health of U.S. civilian population
- Cross-sectional, nationally representative, 80% response rate
- Annual sample of 35,000 households containing 87,500 individuals

Core components of NHIS questionnaire:

- **Household**: basic demographics, geocodes (restricted access)
- **Family**: demographics, food security, program participation, health status, injuries, healthcare utilization, health insurance
- **Sample adult** (one randomly selected adult per family): psychological distress, selected mental health problem, other aspects of health status, health care services, health behaviors
- **Sample child** (one randomly selected child per family): health status, health care services, health behaviors

NHIS also provides imputed income, covers extra topics

2011–2014 Analytical Sample

We pool linked sample adult–family records, NHIS 2011–14:

- Sample adult is aged **18–64** years (working age)
- Every family member is U.S. citizen
- Income \leq **130% of poverty** (gross income cutoff for SNAP)
- ***N* = 21,520**
- *Note 1:* NHIS started administering food security survey module in 2011
- *Note 2:* E.g., in 98% of cases in 2014, “household” is identical to “family”

Selected sample characteristics (weighted):

Variable	Mean	(Std.Dev.)
SNAP participation (indicator)	0.485	(0.500)
Income-to-poverty ratio	0.689	(0.372)
Child (age < 18) present	0.355	(0.479)
Sample adult’s age (years)	37.05	(14.32)
Sample adult is male	0.436	(0.496)

Food Security Indicators

NHIS includes **10-item** food security survey module:

- Referenced to **last 30 days**
- Includes family- and adult-specific questions (no child questions)

We create two indicators of family's food security status:

- 1) **Food secure**: 1 if raw score ≤ 2 (high or marginal FS)
- 2) **Not very low food secure**: 1 if score ≤ 5 (absence of very low FS)

Selected descriptive statistics (weighted):

Indicator	Mean	(Std.Dev.)
Family is food secure	0.677	(0.468)
Family is not very low food secure	0.831	(0.375)

	SNAP subsample	Non-SNAP subsample	Difference
Food secure	0.574	0.775	-0.201 ^{***}

Indicators of Psychological Distress

NHIS administers 6 questions underlying **Kessler (K-6) psychological distress scale**:

- How frequently in past 30 days one felt sad, nervous, restless, hopeless, that everything was an effort, worthless (5-point Likert scale for answers)
- K-6 is standardized and validated measure of **nonspecific psychological distress** (CDC, 2013)

We follow McMorrow et al. (2016) and create indicators for:

- 1) Sample adult in **severe distress**: 1 if K-6 scale ≥ 13 (max is 24)
- 2) Sample adult in **moderate or severe distress**: 1 if K-6 scale ≥ 8

Selected descriptive statistics (weighted):

Indicator	Mean	(Std.Dev.)
Adult is in severe distress	0.097	(0.296)
Adult is in moderate/severe distress	0.226	(0.418)

Indicators of Mental Health Problems

NHIS asks sample adults about degree of **difficulty** with 12 daily activities (e.g., walking) and what health problem caused this

NHIS also asks whether adults are **limited** in performing 7 activities (e.g., personal care) and what health problem caused this

We create indicators for existence of:

- 1) **Mental health problem causing difficulty** with activities
- 2) **Mental health problem causing limitation** in activities

'Problem' includes depression, anxiety, ADD, bipolar, schizophrenia, etc.

Selected descriptive statistics (weighted):

Indicator	Mean	(Std.Dev.)
Adult has mental health problem causing difficulty	0.069	(0.253)
Adult has mental health problem causing limitation	0.083	(0.275)

Food Security on Subsamples

Prevalence of food security (% , weighted) in subsamples by moderate/severe distress and SNAP participation:

		SNAP participation		(SNAP=Yes,·)
		No	Yes	– (SNAP=No,·)
Moderate or severe distress	No	83.03	65.46	-17.56
	Yes	49.43	39.11	-10.32
(Distress=Yes,·) – (Distress=No,·)		-33.60	-26.35	

Also, distress, mental health problem indicators are **positively associated** with SNAP participation

Motivation for Our Methodology

Simple parametric approach:

$$Y_i = \gamma \cdot S_i + x_i \cdot \beta + \varepsilon_i$$

Outcome Treatment Covariates Error term

Treatment S_i is **binary**. E.g., $S_i = 1$ if i is in distress, 0 if not

If same unobservables affect S_i and Y_i , then $cov(S_i, \varepsilon_i) \neq 0$ and OLS is inconsistent due to **endogeneity**

Measurement error in S_i is **nonclassical** \rightarrow standard IV estimation is inconsistent too

Our **nonparametric bounding** methodology handles endogeneity and misreporting. We also develop methods to handle multiple treatments (not just one binary S_i), e.g., treatment = {in distress, on SNAP}

Basics of Our Approach

Define: $H^* = 1$ if adult is truly in distress, $= 0$ otherwise; H is self-reported measure of H^*

We assess **average treatment effect (ATE)** of distress on food security:

$$ATE(1, 0 | X) = P[Y(H^* = 1) = 1 | X] - P[Y(H^* = 0) = 1 | X]$$

$Y = 1$: family is food secure, $Y = 0$: insecure

$Y(H^* = 1)$ indicates **potential** food security outcome if adult were to be in distress. $Y(H^* = 0)$ denotes potential outcome if adult were not in distress

X specifies subpopulation of interest. Say, families with income $\leq 130\%$ of poverty, containing U.S. citizens, sample adult aged 18–64

Not a regression framework: covariates X are not regressors, no regression error term here, no orthogonality conditions to satisfy

Decomposition Strategy

ATE cannot be point-identified without assumptions **even if** $H^* = H$

We decompose every formula into what is identified and what isn't

Let's simplify notation: $ATE = P[Y(1) = 1] - P[Y(0) = 1]$

Consider decomposition:

$$P[Y(1) = 1] = \underbrace{P[Y(1) = 1 | H^* = 1]}_{\text{identified}} \underbrace{P(H^* = 1)}_{\text{identified}} + \underbrace{P[Y(1) = 1 | H^* = 0]}_{\text{not identified}} \underbrace{P(H^* = 0)}_{\text{identified}}$$

Data cannot identify $P[Y(1) = 1 | H^* = 0]$ because it refers to unobserved **counterfactual**. We only know $P[Y(1) = 1 | H^* = 0] \in [0, 1]$

However, using methods of Manski (1995), we can still find worst-case bounds for $P[Y(1) = 1]$, $P[Y(0) = 1]$, and ATE

Addressing Misreporting

$$\begin{aligned}P[Y(1) = 1] &= P(Y = 1, H^* = 1) + P[Y(1) = 1 | H^* = 0]P(H^* = 0) \\ &= P(Y = 1, H = 1) + \theta_1^- - \theta_1^+ + P[Y(1) = 1 | H^* = 0]P(H^* = 0)\end{aligned}$$

$$\theta_1^- \equiv P(Y = 1, H = 0, H^* = 1), \quad \theta_1^+ \equiv P(Y = 1, H = 1, H^* = 0)$$

Sharp **bounds** on ATE:

$$\begin{aligned}P(Y = 1, H = 1) - P(Y = 1, H = 0) - P^* + 2(\theta_1^- - \theta_1^+) \\ \leq ATE \leq\end{aligned}$$

$$P(Y = 1, H = 1) - P(Y = 1, H = 0) + (1 - P^*) + 2(\theta_1^- - \theta_1^+)$$

where $P^* \equiv P(H^* = 1)$

Tightening Bounds

Without assumptions, ATE bounds are **wide** and **contain zero**

To tighten them, we can:

- Use logical constraints on probabilities and auxiliary (validation) data to restrict θ 's
- Apply “**no false positives**” assumption $\rightarrow \theta_1^+ = \theta_0^+ = 0$
- Impose restrictions on selection process:
 - Monotone treatment selection (**MTS**)
 - Monotone instrumental variable (**MIV**)
 - Monotone treatment response (**MTR**)

By layering progressively stronger assumptions we show how they shape inference

Monotonicity Assumptions

Monotone treatment selection (MTS):

$$P[Y(j) = 1 | H^* = 1] \leq P[Y(j) = 1 | H^* = 0], j = 0, 1$$

Monotone instrumental variable (MIV):

Let v be income-to-poverty ratio. Higher v would not harm food security:

$$u_1 \leq u \leq u_2 \Rightarrow$$

$$P[Y(j) = 1 | v = u_1] \leq P[Y(j) = 1 | v = u] \leq P[Y(j) = 1 | v = u_2]$$

Monotone treatment response (MTR):

Poor mental health would not improve food security on average:

$$P[Y(1) = 1 | H^* = h] \leq P[Y(0) = 1 | H^* = h], h = 0, 1$$

Bounds under Endogenous Selection

		Self-reported prevalence rate: $P^* = P = 0.235$			10% Underreporting of true prevalence rate: $P^* = 1.1P = 0.258$		
Endogenous selection		LB	UB	width	LB	UB	width
(a) <u>Arbitrary errors</u>	p.e. [†]	[-0.912,	0.558]	1.469	[-0.935,	0.581]	1.516
	CI [‡]	[-0.919	0.567]		[-0.942	0.590]	
(b) <u>No false positives</u>	p.e.	[-0.710,	0.290]	1.000	[-0.734,	0.313]	1.047
	CI	[-0.716	0.296]		[-0.739	0.319]	

[†] Point estimates of the population bounds.

[‡] Imbens-Manski 95% confidence intervals around the true ATE.

Bounds under MTS+MIV+MTR

		Self-reported prevalence rate: $P^* = P = 0.235$			10% Underreporting of true prevalence rate: $P^* = 1.1P = 0.258$		
MTS + Income MIV + MTR:		LB	UB	width	LB	UB	width
(a) Arbitrary errors	p.e.	[-0.851, -0.0956]		0.756	[-0.878, -0.0956]		0.783
	CI	[-0.878 -0.0649]			[-0.901 -0.0649]		
	bias	+0.011	-0.015		+0.016	-0.015	
(b) No false positives	p.e.	[-0.210, -0.0956]		0.115	[-0.290, -0.0956]		0.194
	CI	[-0.266 -0.0649]			[-0.331 -0.0649]		
	bias	+0.020	-0.015		+0.020	-0.015	

Strictly positive average treatment effects in **bold**.

† Point estimates of the population bounds corrected for finite sample bias.

‡ Imbens-Manski 95% confidence intervals around the true ATE.

* Estimated finite sample bias prior to correction.

Bounds by Reported SNAP Status

Self-reported
prevalence rate:
 $P^* = P$

10% Underreporting
of true prevalence rate:
 $P^* = 1.1P$

SNAP participants (N = 10,918), MTS + MIV + MTR:

		LB	UB	width	LB	UB	width
(a) Arbitrary errors	p.e.	[-0.788,	-0.0779]	0.710	[-0.824,	-0.0779]	0.747
	CI	[-0.820	-0.0364]		[-0.857	-0.0364]	
(b) No false positives	p.e.	[-0.151,	-0.0779]	0.073	[-0.228,	-0.0779]	0.150
	CI	[-0.228	-0.0364]		[-0.302	-0.0364]	

Nonparticipants (N = 10,036), MTS + MIV + MTR:

		LB	UB	width	LB	UB	width
(a) Arbitrary errors	p.e.	[-0.921,	-0.0904]	0.831	[-0.940,	-0.0904]	0.850
	CI	[-0.934	-0.0516]		[-0.953	-0.0516]	
(b) No false positives	p.e.	[-0.251,	-0.0904]	0.161	[-0.336,	-0.0904]	0.245
	CI	[-0.319	-0.0516]		[-0.387	-0.0516]	

Strictly negative average treatment effects in **bold**.

† Point estimates of the population bounds corrected for finite sample bias.

‡ Imbens-Manski 95% confidence intervals around the true ATE.

Next Steps

We have fully developed methods to assess causal effects of mental health on food security and potentially mitigating role of SNAP, while allowing for **both mental health and SNAP to be endogenous and misreported**. Application is underway

To date, we only used income-to-poverty variable as MIV

Next, we will estimate causal effects and use location-specific variables as MIVs to narrow bounds. These include:

- Food store density
- Generosity of food bank and pantry system

Estimations using location-specific information will be performed within an RDC environment

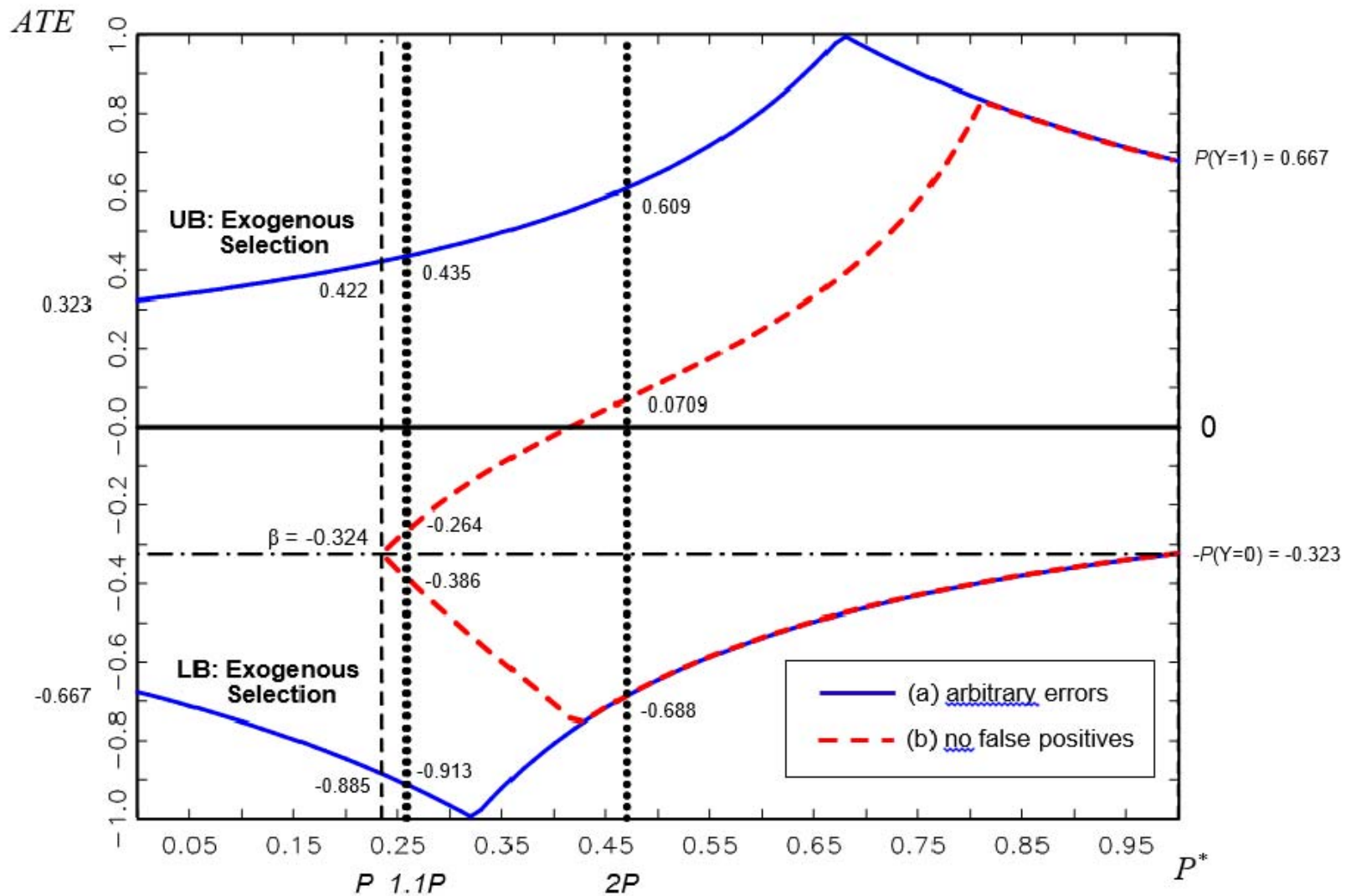


Thank you!



Appendix

Bounds under Exogenous Selection (I)



Bounds under Exogenous Selection (II)

		Self-reported prevalence rate: $P^* = P = 0.235$			10% Underreporting of true prevalence rate: $P^* = 1.1P = 0.258$		
Exogenous selection		LB	UB	width	LB	UB	width
(a) <u>Arbitrary errors</u>	p.e. [†]	[-0.885, 0.422]		1.307	[-0.913, 0.435]		1.348
	CI [‡]	-0.894	0.431		-0.922	0.445	
(b) <u>No false positives</u>	p.e.	[- 0.324 , - 0.324]		0.000	[-0.386, -0.264]		0.123
	CI	-0.341	-0.308		-0.400	-0.250	

Strictly negative average treatment effects in **bold**.

[†] Point estimates of the population bounds.

[‡] Imbens-Manski 95% confidence intervals around the true ATE.

Bounds under Endog. Selection: Graph

